

# Tools for sparse Bayesian deep learning

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# Introduction

1. deep learning (DL) models have tremendous approximation power. But estimation (training) requires lot of data.
2. In data-poor areas, domain knowledge and sparsity may help.

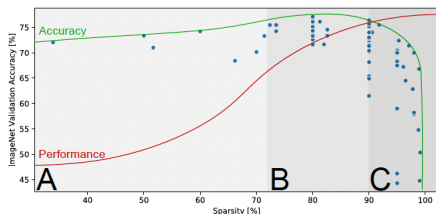


Fig. 4. Typical test error vs. sparsity showing Occam's hill (network: ResNet-50 on Top-1 ImageNet).

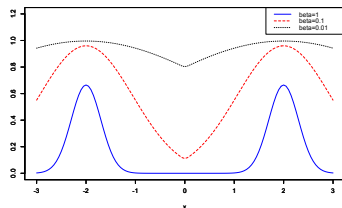
3. The talk discusses two ideas towards that goal: Cyclical MCMC and asynchronous MCMC.

Cyclical MCMC

Asynchronous MCMC

Experimentation with deep learning models

# Cyclical MCMC



1. 'Annealing' / 'tempering'. Let  $\mathcal{E} : \mathcal{X} \rightarrow \mathbb{R}$  with minimum set  $\mathcal{M}$ . Set

$$\pi_t(x) \propto \exp(-\beta_t \mathcal{E}(x)), \quad \beta_t > 0.$$

2. As  $\beta_t \uparrow \infty$ ,  $\pi_t(\cdot) \approx \pi_\infty(\cdot) = \frac{|\cdot \cap \mathcal{M}|}{|\mathcal{M}|}$ .

1. Combined with the Metropolis algorithm and we get Simulated Annealing (SA)

**Thermodynamical Approach to the Traveling Salesman Problem: An Efficient Simulation Algorithm<sup>1</sup>**

V. ČERNÝ<sup>2</sup>

Communicated by S. E. Dreyfus

**Optimization by Simulated Annealing**

[S. KIRKPATRICK, C. D. GELATT, JR., AND M. P. VECCHI](#) [Authors Info & Affiliations](#)

SCIENCE • 13 May 1983 • Vol 220, Issue 4598 • pp. 671-680 • DOI:10.1126/science.220.4598.671

2. For well-designed nonhom. Markov chain  $\{X_t, t \geq 0\}$  with kernels  $\{P_t, t \geq 0\}$  with  $\pi_t P_t = \pi_t$ , and well-chosen sequence  $\beta_t$ ,

$$\mathbb{P}(X_t \in \cdot) - \pi_t(\cdot) \approx 0.$$

# Cyclical MCMC

1. It became quickly clear to the MCMC pioneers that the idea behind SA can be used also to sample from a distribution of interest  $\pi$  by annealing up to 1.
2. Led to parallel tempering (PT) that targets

$$\bar{\pi}(x_1, \dots, x_K) \propto \prod_{k=1}^K \pi(x_k)^{\beta_k}.$$

3. And simulated tempering (ST) that targets

$$\pi(k, x) \propto \exp(-\beta_k \mathcal{E}(x)) / c_k.$$

Simulated Tempering: A New Monte Carlo Scheme

E. Marinari<sup>1,2,3</sup> and G. Parisi<sup>1,2</sup>

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[Europhysics Letters](#), Volume 19, Number 6

Citation E. Marinari and G. Parisi 1992 *EPL* 19 451

DOI 10.1209/0295-5075/19/6/002

Annealing Markov Chain Monte Carlo With  
Applications to Ancestral Inference

Charles J. Geyer and Elizabeth A. Thompson\*

# Cyclical MCMC

1. Unlike SA which remains a mysterious metaheuristics with some theoretical backing, PT and ST benefits from the rigor of MC theory.
2. However these algorithm come with a higher computational price. Costly to use for DL.
3. With Cyclical MCMC, we go back to the original SA framework.

## CYCLICAL STOCHASTIC GRADIENT MCMC FOR BAYESIAN DEEP LEARNING

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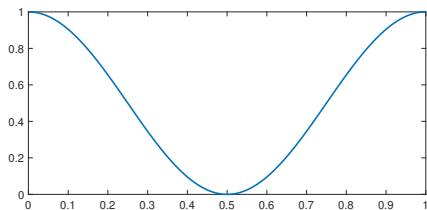
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# Cyclical MCMC

1. Let  $\beta : [0, 1] \rightarrow \mathbb{R}$  such that  $\beta_0 = \beta_1 = 1$ ,  $\beta_t \searrow \nearrow$ .



2. We extend  $t \mapsto \beta_t$  to  $\mathbb{R} \rightarrow \mathbb{R}$  by periodic extension.
3. Let  $\pi(x) \propto e^{-\mathcal{E}(x)}$  a density of interest. For  $k \geq 0$ , we define

$$\pi_k(x) \propto \exp(-\beta_{(k/L)}\mathcal{E}(x)).$$

4. Cyclical:  $\pi_{k+jL} = \pi_k$ .



# Cyclical MCMC

1. Let  $P_k$  be a Markov kernel with invariant distribution  $\pi_k$ .
2. The Cyclical MCMC sampler is a nonhomog. Markov chain  $\{X_k, k \geq 0\}$  with sequence of transition kernels  $\{P_k, k \geq 1\}$ .
3. We collect samples at times  $jL, j = 0, 1, \dots$

# Cyclical MCMC

1. The Cyclical MCMC sampler is a nonhomog. Markov chain  $\{X_k, k \geq 0\}$  with sequence of transition kernels  $\{P_k, k \geq 1\}$ .
2. By periodicity, its can also be viewed as a homogeneous MC  $\{X_{jL}, j \geq 0\}$  with transition kernel

$$P_1 \times \cdots \times P_L.$$

3. Intuition: for well-chosen  $\beta$ ,  $K$  has very good mixing: for  $1 \leq \ell \leq L$ :

$$\{P_1 \times \cdots \times P_\ell\}(x, \cdot) - \pi_\ell(\cdot) \approx 0.$$

4. Existing results towards that includes Holley & Stroock (1991), Douc et al. (2004), Narayanan & Rakhlin (2017), Andrieu et al. (2018).
5. Computationally the algorithm is very efficient.

## Cyclical MCMC: illustration

1.

$$\pi(x) = \frac{1}{25} \sum_{i=1}^{25} \mathcal{N}(x|\mu_i, \Sigma).$$

2. We compare MaLa, cyclical MaLa, SGLD, cyclical SGLD.

3. 150,000 total iterations split into 300 cycles. Collect samples at end of cycles.

sampler	MALA	cMALA	ULA	cULA
SD	$29.52 \pm 6.89$	$4.75 \pm 0.54$	$29.11 \pm 7.44$	$4.57 \pm 0.34$

Table: standard deviation of number of samples in each mode

# Cyclical MCMC: illustration

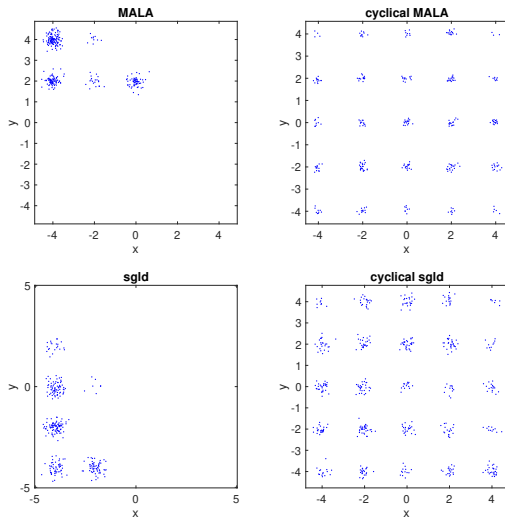


Figure: scatter plots of different method in 25 gaussian mixtures

# Cyclical MCMC

1. On-going work. The cosine cycles works well. But cycle lengths requires careful tuning.
2. We need more theory.

Cyclical MCMC

Asynchronous MCMC

Experimentation with deep learning models

# Asynchronous MCMC for Bayesian sparse deep learning

- ▶ The Gibbs sampler is a hallmark of MCMC methods.
- ▶ A density  $\pi(x_1, x_2)$  on  $X = X_1 \times X_2$ .
- ▶ Let  $\pi_1(\cdot|x_2)$  and  $\pi_2(\cdot|x_1)$  the two conditional distributions.

## Algorithm (Gibbs Sampler)

1. At the  $k$ -th iteration, given  $X^{(k)} = (X_1^{(k)}, X_2^{(k)}) = (x_1, x_2)$ .
    - 1.1 Draw  $\bar{X}_1 \sim \pi_1(\cdot|x_2)$ , and then draw  $\bar{X}_2 \sim \pi_2(\cdot|\bar{X}_1)$ .
  2. Set  $X^{(k+1)} = (\bar{X}_1, \bar{X}_2)$ .
- ▶ Asynchronous Gibbs sampler is a modification of the Gibbs sampler where new random draws are not automatically broadcast.

# Asynchronous MCMC for Bayesian sparse deep learning

- ▶ Asynchronous MCMC was first introduced to the best of my knowledge in the 80's in the CS community as a way of speeding up simulated annealing.
- ▶ Resurfaced again recently in machine learning
  1. Smola and Narayanamurthy (2010) An architecture for parallel topic models. *Proc. VLDB Endow.*
  2. De Sa et al. (2016) Ensuring rapid mixing and low bias for asynchronous gibbs sampling. *ICML 2016 - Volume 48.*
  3. Terenin and Xing (2018). Technique for proving Asynchronous convergence results for MCMC. NIPS 2017.



# Asynchronous MCMC for Bayesian sparse deep learning

- ▶ Asynch. Gibbs sampling does not maintain the correct invariant distribution.
- ▶ For  $a \in [0, 1]$ , suppose that  $X = \{0, 1\} \times \{0, 1\}$ , and

$$\pi(0,0) = 0, \quad \pi(0,1) = \pi(1,0) = \frac{1-a}{2}, \quad \pi(1,1) = a.$$

	0	1
0	0	$(1-a)/2$
1	$(1-a)/2$	$a$

- ▶ If  $\tilde{X}^{(k)} = (1, 1)$ ,

$$\mathbb{P}\left(X^{(k+1)} = (0,0) \mid X^{(k)} = (1,1)\right) = \left(\frac{1-a}{1+a}\right)^2,$$

which will produce a biased sampling asymptotically.

# Asynchronous MCMC for Bayesian sparse deep learning

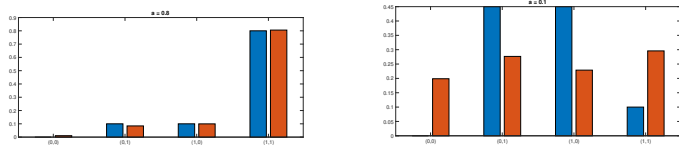


Figure: Gibbs sampler versus asynchronous Gibbs sampler for  $a = 0.8$  and  $a = 0.1$ .

- ▶ The bias is essentially  $\|\pi_{1|2}(\cdot|1) - \pi_{1|2}(\cdot|0)\|_{\text{TV}} = (1 - a)/(1 + a)$ .
- ▶ De Sa et al. (2016) formalized this using Dobrushin coefficient.

# Asynchronous MCMC for Bayesian sparse deep learning

- ▶ Suppose we have a log-likelihood function

$$\ell(\theta) = \ell(\theta, \mathcal{D}) = \sum_{i=1}^n f_{\theta}(z_i), \theta \in \mathbb{R}^p.$$

- ▶ We use a spike and slab prior for  $\theta$ : for  $u > 1$ ,  $0 < \rho_1 < \rho_0 < \infty$ :

$$\delta_j \sim \mathbf{Ber}(p^{-u}), \quad \theta_j | \delta \stackrel{d}{=} \theta_j | \delta_j \stackrel{\text{ind}}{\sim} \begin{cases} \mathbf{N}(0, \rho_1^{-1}) & \text{if } \delta_j = 1 \\ \mathbf{N}(0, \rho_0^{-1}) & \text{if } \delta_j = 0 \end{cases}$$

# Asynchronous MCMC for Bayesian sparse deep learning

The posterior distribution can be written as

$$\Pi(\delta, \theta | \mathcal{D}) \propto \left( p^u \sqrt{\frac{\rho_1}{\rho_0}} \right)^{-\|\delta\|_0} \exp \left( -\frac{\rho_0}{2} \|\theta - \theta_\delta\|_2^2 - \frac{\rho_1}{2} \|\theta_\delta\|_2^2 + \ell(\theta_\delta) \right)$$

- ▶ Asynchronous MCMC algorithm for  $\Pi$ :
  1. fix  $\delta$  and update  $\theta$  (using SGLD or standard MCMC update);
  2. fix  $\theta$  and update  $J$  components of  $\delta$  (using asynchronous Gibbs).

# Asynchronous MCMC for Bayesian sparse deep learning

Why should asynchronous update work here?

- ▶ We have

$$\Pi_j(\delta_j | \delta_{-j}, \theta, \mathcal{D}) \sim \mathbf{Ber}(q_j),$$

where  $q_j$  is driven mainly by  $u \log(p)$  and the log-likelihood ratio

$$\ell(\theta_{\delta^{(j,0)}}) - \ell(\theta_{\delta^{(j,1)}}) \approx -\theta_j \nabla_j \ell(\theta_{\delta^{(j,0)}}) - \frac{\theta_j^2}{2} \nabla_{jj}^{(2)} \ell(\bar{\theta}).$$

- ▶ To illustrate, assume logistic regression.

$$\begin{aligned} \nabla_j \ell(\theta) &= \sum_{i=1}^n \left( Y_i - \frac{e^{\langle \theta, \mathbf{x}_i \rangle}}{1 + e^{\langle \theta, \mathbf{x}_i \rangle}} \right) \mathbf{x}_{ij} = \sum_{i=1}^n \left( Y_i - \frac{e^{\langle \theta_*, \mathbf{x}_i \rangle}}{1 + e^{\langle \theta_*, \mathbf{x}_i \rangle}} \right) \mathbf{x}_{ij} \\ &\quad - (\theta_j - \theta_{*j}) \sum_{i=1}^n D_i(\bar{\theta}) \mathbf{x}_{ij}^2 - \sum_{k \neq j} (\theta_k - \theta_{*k}) \sum_{i=1}^n D_i(\bar{\theta}) \mathbf{x}_{ij} \mathbf{x}_{ik}. \end{aligned}$$

# Asynchronous MCMC for Bayesian sparse deep learning

## Algorithm (Asynch. Sparse SGLD (AS-SGLD))

1. fix  $\delta$  and update  $\theta$  (using Stochastic Gradient Langevin dynamics – SGLD);
2. Given  $\theta$ , select  $J$  components, for  $\vartheta$ , and compute  $G = \nabla \ell(\theta_{\vartheta})$ . Draw independently

$$\delta_{J_k} \sim \mathbf{Ber}(q_{J_k}), \quad q_{J_k} = \left( 1 + \frac{e^{a_0(\theta_{J_k})}}{e^{a_1(\theta_{J_k})}} e^{-\theta_{J_k} G_{J_k}} \right)^{-1}.$$

# Approximate correctness for linear regression

$$\ell(\theta) = -\frac{1}{2\sigma^2} \|y - X\theta\|_2^2, \quad \theta \in \mathbb{R}^p, \quad \sigma^2 > 0 \text{ known.} \quad (1)$$

## Theorem

*Under classical high dim. lin. regr. assumptions and*

$$n \gtrsim \max(\underline{\theta}_*^{-2}(1 + s_*^3) \log(p), J^2 \log(p), (\log(p))^3), \quad \text{and } u \geq C_2(1 + s_*)^2, \quad (2)$$

$$\begin{aligned} \mathbb{E}_* \left[ \max_{j: \delta_{*j}=1} \left| \mathbb{P}(\delta_j^{(k)} = 1) - \Pi(\delta_j = 1 | \mathcal{D}) \right| \right] \\ \leq \left( 1 - \frac{3J}{10p} \right)^k + \exp \left( -C_3 \underline{\theta}_* \sqrt{n} + C_4 J \sqrt{\log(p)} \right) + \frac{10}{p}. \end{aligned}$$

*with probability at least  $1 - 10/p$  (over the data).*

# Logistic regression

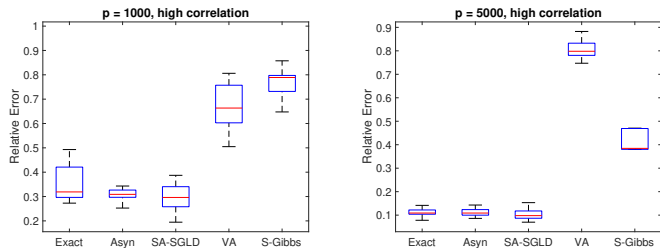


Figure: Relative error for logistic regression model. Based on 50 replications.



# Logistic regression

p/n	Complexity/iteration	1000/500	2000/1000	5000/2500
Exact	$O(nJ\ \delta^{(k)}\ _0)$	5.25s	35.13s	1360.09s
Asyn	$O(n(\ \delta^{(k)}\ _0 + J))$	0.71s	2.19s	99.04s
SA-SGLD	$O(B(\ \delta^{(k)}\ _0 + J))$	0.24s	1.44s	30.12s
Skinny-Gibbs	$O(n(p \vee \ \delta^{(k)}\ _0^2))$	10.50s	87.27s	1154.40s
VA	$O(B \cdot J \cdot p)$	4.05s	34.42s	1243.82s

Table: Running times to convergence

Cyclical MCMC

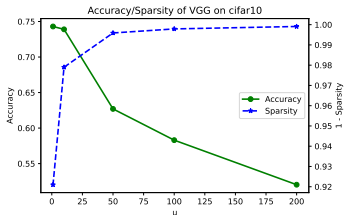
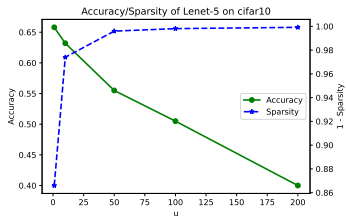
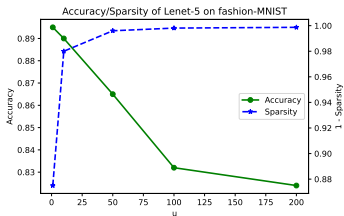
Asynchronous MCMC

Experimentation with deep learning models

# Experimentation with deep learning models

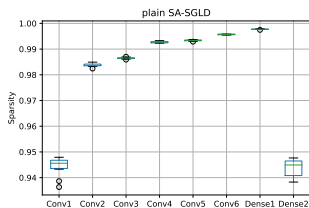
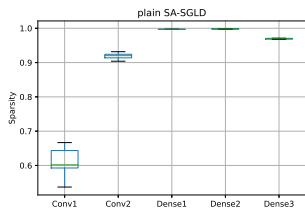
- ▶ Lenet-5 and a baby VGG-16 architectures.
- ▶ Using MNIST-FASHION and Cifar-10 datasets.
- ▶ The goal is to classify small images.

# Experimentation with deep learning models



# Experimentation with deep learning models

## Sparsity of each layer in Lenet-5(left) and VGG(right)



# Experimentation with deep learning models

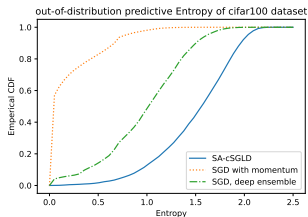
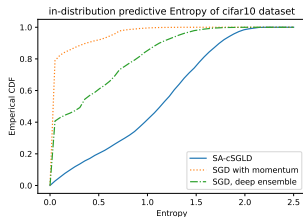
	Accuracy	Density
SGD with Momentum	0.764	1
SGLD	0.8029	1
cSGLD	0.8042	1
plain SA-SGLD, $u = 50$	0.727	0.0047
SA-cSGLD, $u = 50$	0.758	0.0065
SA-SGLD, 10 chains, $u = 50$	0.745	0.0058

Table: VGG-6 with Cifar-10 dataset

# Experimentation with deep learning models

We compare

$$\text{Ent} (p_{\widehat{W}}(\cdot|\mathbf{x})) , \text{ and } \text{Ent} \left( \int p_W(\cdot|\mathbf{x}) \Pi(dW|\mathcal{D}) \right).$$



## Concluding thoughts

- ▶ We have presented two approximate MCMC ideas that we have found very useful for large scale sparse Bayesian modeling.
- ▶ Particularly in low-data and noisy-data settings.
- ▶ More theoretical analysis is needed.
- ▶ In the context of DL, software and hardware to take advantage of sparsity is also needed.

Thanks!!